



Semi-rigorous statistical inference: fitting multiplicities in Au-Au collisions at $\sqrt{s_{NN}} = 62.4$ and 200 GeV



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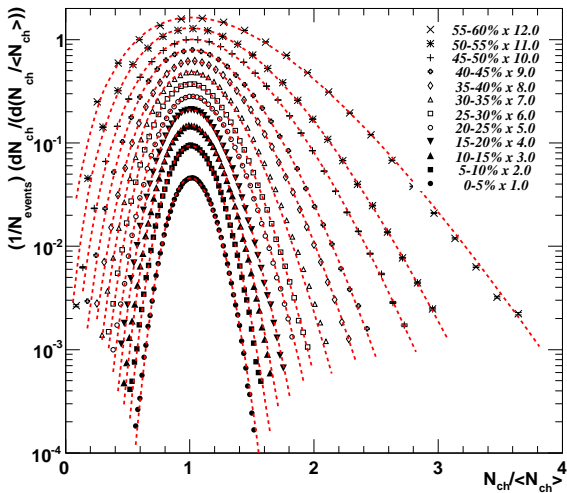
DP, Phys. Rev. C **88**, 034910 (2013)

Data:

PHENIX Collaboration, Phys. Rev. C **78**, 044902 (2008)

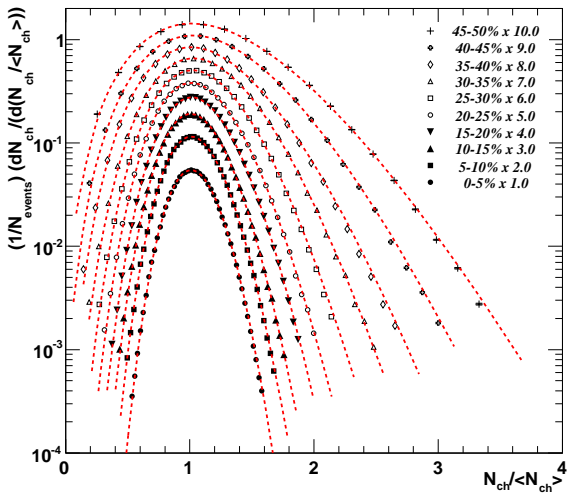


$$|\eta| < 0.26$$





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$$P(n; p, k) = \frac{k(k+1)(k+2)\dots(k+n-1)}{n!} (1-p)^n p^k$$

$0 \leq p \leq 1$, k is a positive real number

$n = 0, 1, 2, \dots$ - the number of charged particles in an event

$$(k, p) \quad \longrightarrow \quad \left(k, \bar{n} = \frac{k(1-p)}{p} \right)$$

\bar{n} - expectation value of n



$$\chi_{LS}^2(\vec{Y}; \vec{\theta}) = \sum_{i=1}^N \frac{[Y_i - \Lambda_i(\vec{\theta})]^2}{err_i^2}$$

$\vec{Y} = (Y_1, Y_2, \dots, Y_N)$ - vector of data

err_i - uncertainty of the i th measurement

$\vec{\theta} = (\theta_1, \theta_2, \dots, \theta_m)$ - parameters of the model

$$\chi_{LS, min}^2(\vec{Y}) = \chi_{LS}^2(\vec{Y}; \hat{\vec{\theta}})$$

$\hat{\vec{\theta}}$ - LS estimators of parameters $\vec{\theta}$



If

- ▶ (Y_1, Y_2, \dots, Y_N) are independent Gaussian random variables with known variances σ_i^2 ,
- ▶ errors $err_i = \sigma_i$,
- ▶ the hypothesis $\Lambda_i(\theta_1, \dots, \theta_m)$ is linear in the parameters θ_j ,
- ▶ the hypothesis is correct,

then the test statistic $\chi_{LS,min}^2$ is distributed according to a χ^2 distribution with $n_d = N - m$ degrees of freedom.

If the hypothesis is nonlinear, then the conclusion is valid in the limit $N \rightarrow \infty$.



$$0 \leq t < +\infty,$$

$n = 1, 2, \dots$ - the number of degrees of freedom

$$f(t; n) = \frac{1}{2^{n/2} \Gamma(n/2)} t^{n/2-1} \cdot e^{-t/2}$$

$$E[t] = n, \quad V[t] = 2n$$



$$E[\chi_{LS,min}^2] = n_d, \quad V[\chi_{LS,min}^2] = 2n_d$$

$$\sigma[\chi_{LS,min}^2] = \sqrt{2n_d}$$

$$\frac{\chi_{LS,min}^2}{n_d} \sim 1$$

$$\frac{\chi_{LS,min}^2}{n_d} = 1 \pm \sqrt{\frac{2}{n_d}}$$



The probability of obtaining the value of the test statistic equal to or greater than the value just obtained for the present data set (*i.e.* χ_{min}^2), when repeating the whole experiment many times (repeating measurement of \vec{Y}):

$$p = P(\chi^2 \geq \chi_{min}^2) = \int_{\chi_{min}^2}^{\infty} g(t) dt ,$$

$g(t)$ - probability density function of χ_{min}^2 , NOT KNOWN USUALLY

$\chi_{min}^2(\vec{Y})$ - statistic because a function of multidimensional random variable \vec{Y}



- ▶ Assume the significance level α in advance.

($\alpha = 0.1\%$, here)

- ▶ If $p < \alpha$, a hypothesis should be rejected ("bad fit").
- ▶ If $p \geq \alpha$, a hypothesis can not be rejected ("good fit").



$$\chi_{LS}^2(\vec{n}; \bar{n}, k) = \sum_{i=1}^m \frac{(n_i - \nu_i(\bar{n}, k))^2}{err_i^2}$$

$\vec{n} = (n_1, n_2, \dots, n_m)$ - vector of data (entries)

err_i - uncertainty of the i th measurement

$\nu_i = N \cdot P(i - 1; \bar{n}, k)$ - expected number of entries

$N = \sum n_i$ - total number of events

$$\chi_{min}^2(\vec{n}) = \chi_{LS}^2(\vec{n}; \hat{n}, \hat{k})$$

\hat{n}, \hat{k} - estimators of parameters \bar{n} and k



$$err_i^2 = \sigma_{i,stat}^2 + \sigma_{i,syst}^2$$

$$\sigma_{i,stat} = \sqrt{n_i}, \quad \sigma_{i,syst} = 3 \cdot \sigma_{i,stat} = 3 \cdot \sqrt{n_i}$$

$$\chi_{PHEN}^2(\vec{n}; \bar{n}, k) = \frac{1}{10} \cdot \sum_{i=1}^m \frac{(n_i - \nu_i(\bar{n}, k))^2}{n_i} = \frac{1}{10} \cdot \chi_N^2(\vec{n}; \bar{n}, k)$$

χ_N^2 - Neyman's χ^2 test statistic, asymptotically χ^2 distributed !
Jerzy Słowia-Neyman (1894-1981)

\implies PHENIX χ^2 function is NOT χ^2 distributed !



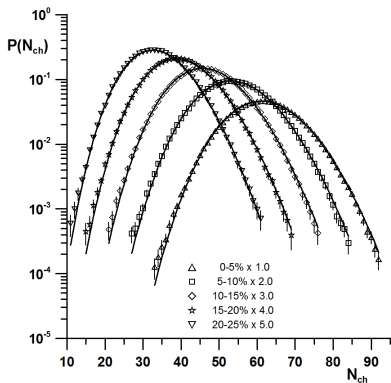
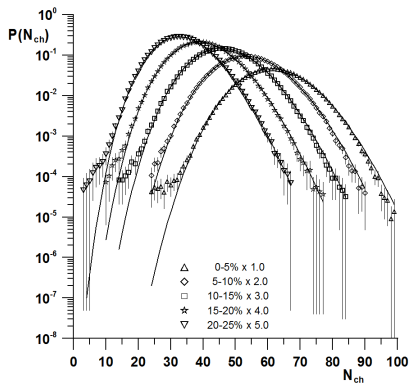
the distribution $g(t)$ of a function $t(z)$ of a random variable z with the known p.d.f. $f(z)$:

$$g(t) = f(z(t)) \left| \frac{dz}{dt} \right|$$

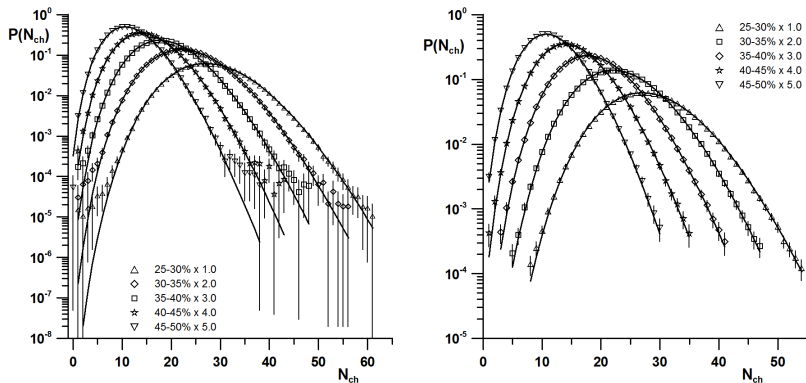
$$g(t; n_{dof}) = 10f(10t; n_{dof})$$

p -value of PHENIX test statistic:

$$p = \int_{10 \cdot \chi_{PHEN, min}^2}^{\infty} f(t; n_{dof}) dt$$



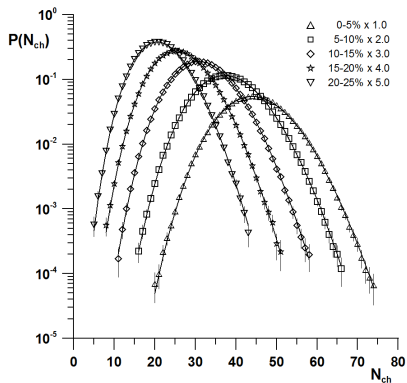
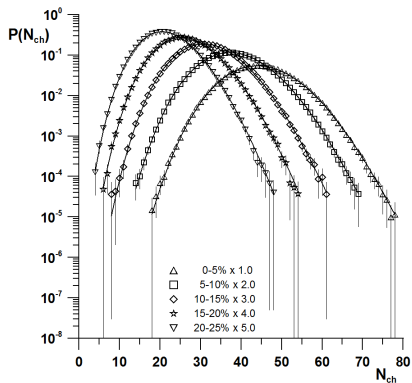
Rysunek: Uncorrected multiplicity distributions for bins with $n_i > 5$ (left) and $n_i > 60$ (right).



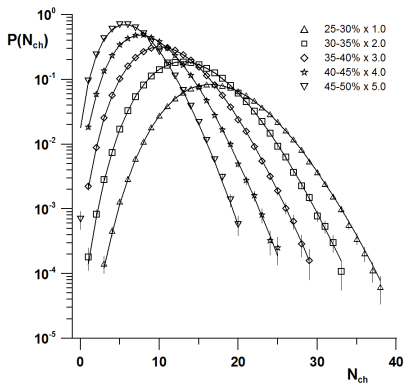
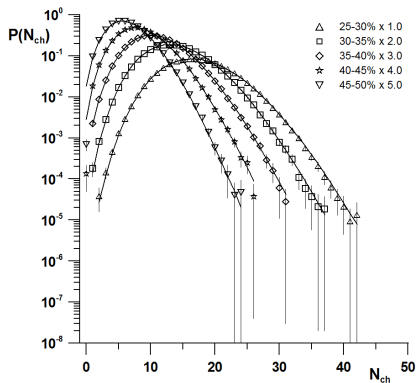
Rysunek: Uncorrected multiplicity distributions for bins with $n_i > 5$ (left) and $n_i > 60$ (right).



Centr. %	N	\hat{k}	\hat{n}	χ^2_{PHEN}/n_d	p -value %
0-5	652579	289.0 ± 2.9	61.86 ± 0.01	0.57	0
5-10	657571	168.1 ± 1.2	53.91 ± 0.01	0.61	0
10-15	658258	116.4 ± 0.7	46.50 ± 0.01	0.53	0
15-20	659302	86.9 ± 0.5	39.72 ± 0.01	0.43	0
20-25	658461	69.1 ± 0.4	33.56 ± 0.01	0.34	0
25-30	659337	57.9 ± 0.3	28.0 ± 0.01	0.28	$6.7 \cdot 10^{-8}$
30-35	659021	48.3 ± 0.3	23.02 ± 0.01	0.16	0.76
35-40	660937	41.3 ± 0.2	18.64 ± 0.01	0.19	0.12
40-45	661422	34.6 ± 0.2	14.84 ± 0.01	0.21	0.015
45-50	661577	27.9 ± 0.2	11.56 ± 0.005	0.23	0.011
50-55	661877	21.9 ± 0.1	8.81 ± 0.004	0.30	$7.8 \cdot 10^{-5}$



Rysunek: Uncorrected multiplicity distributions for bins with $n_i > 5$ (left) and $n_i > 40$ (right).



Rysunek: Uncorrected multiplicity distributions for bins with $n_i > 5$ (left) and $n_i > 60$ (right).



Centr. %	N	\hat{k}	\hat{n}	χ^2_{PHEN}/n_d	p -value %
0-5	607075	227.9 ± 2.5	44.67 ± 0.01	0.19	$5.6 \cdot 10^{-3}$
5-10	752263	143.9 ± 1.1	37.96 ± 0.01	0.12	14.4
10-15	752739	116.2 ± 0.9	31.53 ± 0.01	0.13	7.0
15-20	752492	88.5 ± 0.6	26.07 ± 0.01	0.11	30.9
20-25	752182	69.2 ± 0.5	21.35 ± 0.01	0.22	$2.4 \cdot 10^{-3}$
25-30	752095	53.6 ± 0.4	17.30 ± 0.01	0.23	$1.8 \cdot 10^{-3}$
30-35	751324	40.3 ± 0.3	13.84 ± 0.005	0.26	$4.3 \cdot 10^{-4}$
35-40	751639	31.8 ± 0.2	10.89 ± 0.004	0.15	3.5
40-45	750852	25.2 ± 0.2	8.42 ± 0.004	0.22	0.062
45-50	751348	22.0 ± 0.2	6.41 ± 0.003	343	0



1. Caution is necessary, when one infers about quality of a fit when the distribution of a test statistic is not known. Then inference from the condition $\chi^2/n_d \sim 1$ could be confused.
2. Adding statistical and systematic errors in quadrature could change properties of the LS test statistic entirely.
3. As far as PHENIX Au-Au data are concerned, only for 6 from 21 cases of collision energy and centrality the NBD hypothesis can not be rejected.